

# Will the Global Village Fracture into Tribes?

## Recommender Systems and their Effects on Consumer Fragmentation

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### **Abstract**

Personalization is becoming ubiquitous on the World Wide Web. Such systems use statistical techniques to infer a customer's preferences and recommend content best suited to him (e.g., "Customers who liked this also liked..."). A debate has emerged as to whether personalization has drawbacks. By making the web hyper-specific to our interests, does it fragment internet users, reducing shared experiences and narrowing media consumption? We study whether personalization is in fact fragmenting the online population. Surprisingly, it does not appear to do so in our study. Personalization appears to be a tool that helps users widen their interests, which in turn creates commonality with others. This increase in commonality occurs for two reasons, which we term volume and product mix effects. The volume effect is that consumers simply consume more after personalized recommendations, increasing the chance of having more items in common. The product mix effect is that, conditional on volume, consumers buy a more similar mix of products after recommendations.

## 1. INTRODUCTION

Recommender systems are becoming integral to how consumers discover media. They are used for all major types of media, such as books, movies, music, news, and television. They are commonplace at major online firms, such as Amazon, Netflix, and Apple’s iTunes store. And they have a strong influence on what consumers buy and view. With movies, Netflix reports that over 60% of their rentals originate from recommendations (Thompson 2008). With online news, Google News reports that recommendations increase articles viewed by 38% (Das et al. 2007). At Amazon, which sells music, books, and movies, 35% of sales are reported to originate from recommendations (Lamere & Green 2008).

The value that recommenders offer is personalization: the consumption experience is personalized to each user’s taste. A personalized radio station plays music not for the general public but for each particular user. A personalized newspaper does not show the same front page to everyone but customizes it for each reader. A retailer arranges its online shelves and displays based on who is browsing at that moment. Such personalization is valuable in modern media markets, which can have millions of products to choose from. As a result, personalization has also become a major theme of research in Information Systems (e.g., Murthi and Sarkar 2003; Dellarocas 2003; Brynjolfsson et al. 2006; Clemons et al. 2006) and Marketing (e.g., Ansari et al. 2000; Manchanda et al. 2006; Shaffer and Zhang 1995; Rossi et al. 1996), with its origins in targeted and customized marketing.

The following examples show how recommenders systems create this personalized experience:

The newspaper ... is undergoing the most momentous transformation.... Online versions are proliferating, ... yet so far, few newspaper sites look different from the pulp-and-ink papers that spawned them.... Often, the front page changes only once a day, just like the print version, and it shows the same news to all readers. There's no need for that uniformity. Every time a Web server generates a news page, ... it can generate different front pages, ... producing millions of distinct editions, each one targeting just one person – you.

–Greg Linden, creator of Findory news and Amazon recommendations (2008)

Last.fm connects you with your favorite music and uses your unique taste to find new music, people, and concerts you'll like.

–Last.fm website

Along with the benefits of personalization, however, a debate has emerged as to its broader impact on consumers. Personalizing websites means that we may no longer see the same newspaper articles, television shows, or books as our peers. Some critics thus argue that recommenders systems will create fragmentation, causing users to have less and less in common with one another. An alternative view contends that recommenders may do the opposite: recommenders may have homogenizing effects

because they share information among users who otherwise would not communicate. Fragmentation in consumption has implications for consumers and society. For consumers, shared consumption often has an associated externality. For example, at the proverbial water cooler, people are able to discuss a shared book, artist, or news article. This is a positive externality from shared consumption (Katz and Shapiro 1985). If recommenders affect consumption similarity, these externalities are in turn impacted. Furthermore, a key promise of recommenders is that they can help consumers discover new and relevant items outside their sphere of interest. Understanding whether these systems aid such discovery or instead push us down an echo chamber with narrow consumption helps us better understand whether recommenders are delivering on that promise. From a societal perspective, the literature has expressed concern that fragmentation is a negative consequence for society because social discourse suffers when people have a narrow information base with little in common with one another (Sunstein 2007). This could occur with many forms of personalization but is most relevant for personalized news. These critics suggest the media and government should do more to increase exposure to a variety of content. In contrast, finding evidence of homogenization would suggest that such policies and regulation of personalization on the internet are not warranted. This paper presents the first empirical evidence for the debate on whether recommenders fragment versus homogenize users.<sup>1</sup>

We find, in an empirical study of a music industry recommendation service, that recommendations are associated with an increase in commonality. This increase in purchase similarity occurs for two reasons, which we term volume and product mix effects. The volume effect is that consumers simply purchase more after recommendations, increasing the chance of having purchases in common with others. The product mix effect is that consumers buy a more similar mix of products after recommendations, conditional on volume. When we view consumer purchases as a similarity network before versus after recommendations, we find that the network becomes denser and smaller, or characterized by shorter inter-user distances. We find that this increase in commonality occurs because the system helps users explore and discover new items. These findings suggest that, for this setting, concerns of fragmentation may be misplaced.

We note that our results are derived for one recommendation technology deployed in one setting and it is unclear whether the same results will arise for books, news and other settings. Nonetheless, the results demonstrate that a common criticism that these systems cause fragmentation is not universally true and that commonly used designs can in fact increase commonality in consumption.

## 2. PRIOR WORK

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<sup>1</sup> In trying to understand the impact of recommenders on fragmentation, we note that it is possible to deliberately design systems with the goal of increasing commonality or similarly with the goal of decreasing commonality. However, commonality is not a design goal in practice and instead a side-effect of recommender use. Thus, our goal is to document the impact of a commonly used design rather than to investigate if there exists a design that can increase commonality or cause fragmentation.

A simplified taxonomy of recommender systems divides them into content-based versus collaborative filtering-based systems. Content-based systems use product information (e.g., genre, mood, author) to recommend items similar to those a user rated highly. Collaborative filters, in contrast, are unaware of a product's content and instead use correlations in sales or ratings to identify what similar customers bought or liked. Perhaps the best-known collaborative filter is Amazon.com's, with its tagline, "Customers who bought this also bought..." Content-based recommenders do well when there is rich information about product attributes but cannot recommend across product categories or genres. Collaborative filters can recommend across product or genres and do not require knowledge of product attributes but need sufficient purchase/ratings data for users and items in order to recommend. Hybrid systems combine the best of the two approaches. The design of these systems has been an active research area for almost twenty years. An extensive review in the Information Systems literature is provided in Adomavicius & Tuzhilin (2005).

Although a large body of work exists on designing recommenders systems, we know less about how they affect the market and society. This is despite the large body of work on recommender algorithms and millions of transactions occurring through them. This paper continues a small stream of work in that direction. Recommenders can have a positive effect on sales and web impressions (Ansari et al. 2000; Das et al. 2007; Bodapati 2008; De et al. 2010). For example, De et al. (2010) show that recommenders positively affect sales, and they expect this to be particularly important in industries with many SKUs. The question of how recommenders affect *products* has recently been studied: which products gain versus lose sales due to recommenders and whether recommenders increase the market for niche goods, or "long tail" (Fleder and Hosanagar 2009; Oestreicher-Singer and Sundararajan 2012; Hervas-Drane 2013). It was commonly assumed that recommenders increase the long tail, and we now know this is not always true (Fleder and Hosanagar 2009). This paper asks the complementary question of how recommenders affect *consumers*: whether they cause consumers to consume more or less in common with one another.

A range of views exist as to whether recommenders will fragment versus homogenize users. Sunstein (2007) argues that recommenders create fragmentation by limiting users' media exposures to their predefined, narrow interests. He argues that while "some of the recommendations from Amazon.com and analogous services are miraculously good, ... it might well be disturbing if the consequence is to encourage people to narrow their horizons, or to cater to their existing tastes rather than to form new ones" (2007). These fragmentation effects, he argues, can also have undesirable societal consequences in which people live in echo chambers and cannot relate to the views of others. Pariser (2011) similarly argues that online personalization, which includes recommenders, creates a filter bubble – an invisible, personal universe of information – due to which the world each user sees online may be very different

(Terdiman 2011). “In an age when shared information is the bedrock of shared experience, the filter bubble is a centrifugal force pulling us apart,” he argues (Pariser 2011). Pattie Maes, who created one of the first recommender systems, also believed some recommenders could have a narrow-minded and hyperpersonalized aspect (Thompson 2008).

Sunstein, Pariser and Maes all believe that recommenders can fragment users because, in their opinion, showing media close to one’s interests can limit exploration and discovery in other areas. This can limit diversity in individual consumption and reduce commonality in consumption across users. These authors differ in their views as to why fragmentation is undesirable. Sunstein argues that society benefits when people have a range of viewpoints. For example, in news programming, users should be exposed to multiple views on a topic, not just the one that reinforces their existing beliefs. Or as technologist Lanier puts it, “People tend to get into this echo chamber where more and more of what they see conforms to the idea of who some software thinks they are... You start to become more and more like the image of you because that is what you are seeing” (Lanier 2010; Singer 2011). Pariser adds a different critique: fragmentation would cause poor decision making because “the filter bubble confines us to our information neighborhood, unable to see or explore the rest of the enormous world of possibilities” (Pariser 2011). A final critique of fragmentation is lost externalities: a product’s popularity can have a positive externality. The benefit of seeing the same movies as others is that we can discuss them. Under fragmentation, this benefit would disappear. As Maes says, “You don’t want to see a movie just because you think it’s going to be good. It’s also because everyone [else is] ... talking about it, and you want to be able to talk about it too.” Consuming the same media and products is a way of participating in society (Thompson 2008).

We agree with these critics that excessive fragmentation could be undesirable. However we believe that the antecedent, that recommenders create fragmentation, is ultimately an assumption. This paper tests that assumption. If recommenders do not create fragmentation, the proliferation of this view could be harmful to the adoption of otherwise valuable technologies. Part of the promise of recommenders is that they can help us discover new items outside our comfort zone and thereby expand our horizons. We believe that their view of the role of recommenders is rather narrow and a more moderate view is appropriate.

Others have offered moderate views. Nicholas Negroponte, co-founder of the MIT Media Lab, coined the term “The Daily Me” (1995), referring recommenders creating newspapers customized for each person’s interests. The Daily Me might create fragmentation by showing users only the content that matches their viewpoints. However, Negroponte also coined the “The Daily Us,” suggesting that consumers may also turn to recommenders when they need help exploring areas outside their interests. Van Alstyne and Brynjolfsson(2005) formalize this mixed view in an economic model. They ask whether

internet technologies like recommender systems will lead to fragmentation versus homogenization – in their terms, a cyber-Balkans versus a global village. In their model, as technology lowers search costs and communication costs, either outcome can occur. The outcome depends on a parameter representing consumers' preference for specialization. This parameter is difficult to specify, and so complementary empirical work is needed.

Similar mixed views were shared by the creators of early recommenders. When research on recommender systems was just beginning, Paul Resnick, asked if the global village would fracture into tribes (Arnheim 1996). John Riedl, co-inventor of one of the first recommenders, asked if collaborative filtering would democratize information or result in social fragmentation. Greg Linden, who helped develop Amazon's original recommender system, states that “[critics] talk about personalization as narrowing and filtering.... That is not what personalization does. Personalization seeks to enhance discovery” (Linden 2011).

The discussion reveals that there are mixed views as to whether recommenders will fragment users, but there is not yet any empirical evidence on the issue. The goal of this study is to provide the first empirical evidence on the impact of recommenders on purchase similarity.

### **3. PROBLEM FORMULATION**

While many authors have discussed the fragmentation question qualitatively, the empirical question has not been posed in concrete terms. This section defines the problem formally.

#### **3.1. Research Questions**

Throughout this paper, we operationalize the notion of fragmentation and homogenization in terms of commonality in items consumed by users. This is analogous to Van Alstyne and Brynjolfsson's analysis of how the Internet affects knowledge overlap among users (2005) and is consistent with notions of fragmentation as suggested by Sunstein, Pariser and Maes (“shared experiences”). We do not focus on underlying preferences of users and whether recommenders cause them to converge or diverge. Instead, consistent with the main thrust of the fragmentation debate, we focus exclusively on overlap in the items consumed. Changes in this overlap may arise due to changes in preferences or changes in users' product awareness due to recommendations.

Our goal is to study whether recommenders make users' consumption more or less similar to one another. We divide the question in two components:

1. Aggregate level: overall, are consumers farther or closer to one another?
2. Disaggregate level: are there differential effects at the individual level, by which some users become closer and others farther?

The first question measures the overall effect of whether users become farther or closer to one

another. The second question explains why. For example, effect (1) may show that users are less similar on average. Effect (2) explains why: for example, even though there is a net reduction in purchase similarity, it may be the case that the closest users became closer and the farthest became much farther.

A note on terminology: The meaning of “close” and “far” will be quantified in the next section. Qualitatively, throughout the paper we refer interchangeably to users who are “close” as exhibiting similarity, commonality, or homogeneity; opposite this, we refer to users who are “far” as exhibiting fragmentation and having little overlap in their purchases.

### 3.2. Two Group Design

The analysis design throughout is analogous to a two-group experiment. One group is “treated” with recommendations and their behavior compared before versus after. A control group is not treated with recommendations, and their behavior is compared over the same period. The data are in fact observational, as we will discuss, but the terminology of experiments simplifies the writing.

Let  $O_{it}$  denote an observation on group  $i$  during time period  $t$ .  $O_{it}$  is a set of tuples. For our music data, a tuple is of the form {user, artist, # songs purchased} for all users in group  $i$  during period  $t$ . Group  $i = 1$  is the treated group, which is unexposed to the recommender during  $t = 1$  but exposed to the recommender during  $t = 2$ . Group  $i = 2$  is the control, which is unexposed to the recommender during both time periods. The time periods are the same for both groups. Figure 1 represents this setup, where  $X$  denotes exposure to recommendations.

Treated:	$O_{11}$	$X$	$O_{12}$
Control:	$O_{21}$		$O_{22}$

**Figure 1.** Schematic of the Two Group Design

Using this design, we can compare the treated group before and after recommendations. We can also compare the treated group to the control over the same period. The control accounts for factors such as time trends and maturation that might be confounded with recommender usage in a one group pre-post design (Campbell & Stanley 1963).

### 3.3. Hypotheses to Test

We wish to compare how the treated and control groups change over time. Let  $T(O_{it})$  be some statistic of interest on  $O_{it}$  measuring fragmentation. As shorthand, we will write  $T_{it}$ . We define the following quantities of interest:

Difference in treated:	$D_1$	$\equiv T_{12} - T_{11}$
Difference in control:	$D_2$	$\equiv T_{22} - T_{21}$
Difference-in-differences:	$D$	$\equiv D_1 - D_2$

$D_1$  describes changes in the treated group.  $D_2$  describes changes in the control. The difference-in-difference estimator,  $D$ , describes how much changes in the treated group exceed those in the control. For example, suppose that independent of recommendations, a time trend is occurring in the music industry that affects both groups. Thus observing  $D_1 \neq 0$  does not mean recommendations have an effect on consumers because the same trend will affect  $D_2$ . However, the difference-in-differences estimator  $D$  can identify changes in the treated group beyond the time trend by subtracting the change in the control.

Let  $\mu \equiv E[D]$ , where  $D$ 's distribution is not known to us. The central questions of this paper take the form

$$H_0: \mu \equiv E[D] = 0$$

$$H_a: \mu \equiv E[D] \neq 0$$

The above formulation is general for any underlying  $T$ , and many questions about similarity in consumer purchases can be posed in this framework. Several statistics of interest  $T(\cdot)$  are defined in the next section. Each gives rise to a separate  $D$  and hence a separate hypothesis of the form above. The hypotheses are always stated as two-sided. This makes our tests more conservative, but it is necessary because the literature offers mixed views as to whether fragmentation versus homogenization will occur.

## 4. FORMULATION SPECIFICS

This section defines the quantities of interest  $T(O_{it})$ . To facilitate this, we take the intermediate step of defining a network  $G(O_{it})$  among the firm's consumers and making  $T(G(O_{it}))$  a function of that network. At first glance, introducing networks appears to complicate the analysis by adding an extra step. In contrast, we will see this provides a great service for interpreting the data.

### 4.1. Motivation for Network Analysis

We define a network in which consumers are the nodes and edges represent similarity between consumers' purchases. This paper's goal of asking whether users' purchases become more or less similar after recommendations will become equivalent to asking how the consumer network changes pre-post recommendations.

For each  $O_{it}$  we will create a user network  $G(O_{it})$ . Then, we will define quantities of interest (e.g., median degree, average distance) on the network  $T(G(O_{it}))$  and study how these quantities change before versus after recommendations.

The consumer network is not a true social network because its edges do not represent physical relationships. Its edges instead represent similarity in purchases. Still, we find it useful to formulate the problem as a network one. First, the benefit of introducing networks is interpretation. Networks are a useful object for describing changes in user similarity. It is easy to conceive of a network expanding, shrinking, or becoming more dense. In contrast, such interpretations would be difficult if we instead studied a large correlation matrix of users' purchases. Second, network analysis is recently being applied to settings like ours in which edges represent similarity of purchases. Huang et al. (2007) and Smith et al. (2007) use co-purchase and co-occurrence data to build an "implicit" network of individuals. In these examples, the network is not strictly necessary for measuring similarity, but it aids in interpretation.

#### 4.2. Defining the Network

Mathematically, our network is a graph made of nodes and edges. Users are the nodes, and edge weights describe the similarity between user pairs, as defined by commonality in purchases.

For notation, we can interpret  $O_{it}$  as a users  $\times$  artists matrix of purchase counts. An element  $(O_{it})_{xa}$  is the number of songs user  $x$  purchased of artist  $a$ .<sup>2</sup> A row of this matrix is denoted  $(O_{it})_x$ . For each  $O_{it}$ , the corresponding network is  $G(O_{it})$  which is denoted as simply  $G_{it}$ .<sup>3</sup> The network  $G_{it}$  is a users  $\times$  users matrix of edge weights. An element  $(G_{it})_{xy}$  is the edge weight between user  $x$  and user  $y$ . Defining the network is thus equivalent to defining the distance between any two users.

Our main network uses a weighted network construction. Within a given group and time period, users  $x$  and  $y$  have an edge between them with weight given by the cosine distance.

$$\begin{aligned} \text{Cosine Distance Network} \\ (G_{it})_{xy} &\equiv 1 - \text{Cos}((O_{it})_x, (O_{it})_y) \\ &= 1 - \frac{(O_{it})_x \cdot (O_{it})_y}{\|(O_{it})_x\| \|(O_{it})_y\|} \end{aligned}$$

The  $\cdot$  symbol is the vector dot product. The cosine between two users' vectors is a measure of the angle between them. It reflects how similar these users are in the space of artists. Thus 1 minus the cosine is how different they are. This measure is perhaps the most common similarity metric used for analyzing purchase data and the design of recommender systems. Using it in our context, it gives rise to a weighted, undirected network among users.

There are many other ways to construct the network. In weighted networks like the one above, there can be other way to define the edge weights such as Euclidean distance. A simple unweighted

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<sup>2</sup> The vector is defined in terms of artists rather than songs because the recommender used in our study operates at the artist level. That is, the input to the recommender is the artist being played. We discuss the recommender design in detail in Section 5.1.

<sup>3</sup>  $G(\cdot)$  is a function that converts the purchase matrix into a network, or  $G(O_{it}) \equiv G_{it}$ .

network definition is also possible where users  $x$  and  $y$  have an edge between them if they purchase at least one artist in common. In the main sections of this paper, we focus our base case on the cosine-based network above because its definition is simple and intuitive. In the electronic companion, we present results for other network definitions. We simplify the exposition in this way, since all of the networks tested yield nearly the same conclusions.

#### 4.3. Defining $T$ : Measures of the Network's Properties

With the network  $G(O_{it})$  defined, we next define summary statistics of the network's properties,  $T(G(O_{it}))$ .  $T$  summarizes in one number a particular network property and thus facilitates comparisons of the network over time. We define two such measures below. As notation, let  $d_x \circ \hat{a}_{y=1, y^1 x}^n (G_{it})_{xy}$  be the degree of user  $x$  where  $n$  is the number of users in the network. Further, let  ${}_nC_2$  denote the number of user pairs that can be formed from a set of  $n$  users ( ${}_nC_2 = n(n-1)/2$ ).

Measure	$T(G(O_{it})) =$
Median Degree	$\text{Median}\{d_x\}_{x=1}^n$
Average Distance	$\frac{1}{{}_nC_2} \hat{a}_{x=1}^n \hat{a}_{y < x} (G_{it})_{xy}$

*Median Degree.* For a weighted network, the median degree is the sum of distance (edge weights) to other users that the typical (median) user has (Newman 2004).

*Average Distance.* The average distance simply averages the pairwise distance of the users where the distances are measured by cosine distance metric.

To summarize the analysis setup, the data are in the form of a two group experiment ( $O_{it}$ ). Each data set is converted to a network  $G(O_{it})$ . Summary statistics are computed on each network  $T(G(O_{it}))$ . Finally, these statistics are compared across the groups and time.

## 5. DATA

### 5.1. Data Source

We study the fragmentation question using data from an online music service, referred to here as Service. Service is a free software add-on to Apple's iTunes. iTunes, in turn, is the music player that allows users to buy music from Apple's iTunes store, the largest music retailer in the U.S. (Apple 2008). Service personalizes the user experience as follows. When users listen to music in iTunes, Service suggests other songs that the user may like. The suggestions appear in a window appended to iTunes, where the user can sample these songs and opt to purchase them. If a purchase results, Service earns a commission. Service

also provides recommendations through a website where users can view the play histories of other Service users with similar libraries. These play histories are uploaded automatically by the plugin to Service’s website on a continual basis. Together, these two features comprise the personalization technology.<sup>4</sup>

Figure 2 shows a screenshot of the plugin. Apple’s iTunes appears at left. The plugin, as appended to iTunes, is at right and displays a list of recommended songs. The song suggestions by the plugin are based on the artist currently playing (i.e., the query to obtain recommendations is the current artist). Based on the current artist, Service identifies the six most similar artists and populates the window with this list. Artist-to-artist similarity is defined by a hybrid of content and collaborative data, though the results are heavily weighted toward the content portion (90% versus 10%). Thus in the taxonomy of recommender systems by Adomavicius & Tuzhilin (2005), the plugin is for the most part a content-based, item-to-item-based system.



**Figure 2.** Screen shot of the recommendation service.

## 5.2. Novelty of the Data

To study the effects of recommenders, a contrast is needed between users exposed and unexposed to recommendations. The data collected by most retailers (e.g., Amazon, Netflix) is inadequate because retailers only observe consumers after they arrive at their website and hence after exposure to recommendations. This may be the reason, we speculate, that others have not been able to study the fragmentation question. Our data are novel in this regard. When a user registers for Service, a history file is extracted from the user's iTunes player. This history file contains the names and timestamps of all songs ever added to that user's music library, and thus it provides a record of the user's behavior *prior* to

<sup>4</sup> It is common for most online firms to use multi-component recommendation systems. For example, Netflix and Amazon's recommendations page in fact have different types of recommendations generated from different algorithms all on the same page. In these environments, it is hard to isolate the impact of any one component. It is also debatable whether the researcher would prefer to analyze users exposed to just one component of Service's personalization technology. An analysis based on just one component would not be indicative of the real trend occurring online because users are typically exposed to all components. Accordingly, our study focuses on the net impact of the multi-component system.

joining Service. The user's post-registration purchases are also observed by Service because the plugin notifies Service via the internet of all songs added to the user's iTunes library, whether bought at the iTunes store or not. This combination of the history file and continued communication via the plugin thus gives us a before and after view of the user's behavior.

Besides comparing users' purchase histories before and after registering, we can also compare these users with a control group. The control data are obtained by again exploiting the history files of Service users. For users who register after our study, their history files allow us to look backward at their Service-uninfluenced behavior during the same time period. More detail is given in the next section. This use of eventual Service users for the control affords a measure of similarity between the groups. Thus the new data source enables a before-after recommendations contrast as well as data on a control group for the same period.

### **5.3. Data Inclusion Criteria**

This section describes the process for setting up the data in the two-group design introduced earlier. Figure 3 summarizes the details of this process. The data are collected via Service's plugin that is installed on each user's machine. The plugin relays to Service in near real-time the timestamp and product information of any song added to that user's iTunes library. For ease of writing, we refer to songs as purchases, but our data in fact capture all songs added to a user's library, whether purchased from Apple's iTunes store, purchased from another firm, or downloaded elsewhere online.

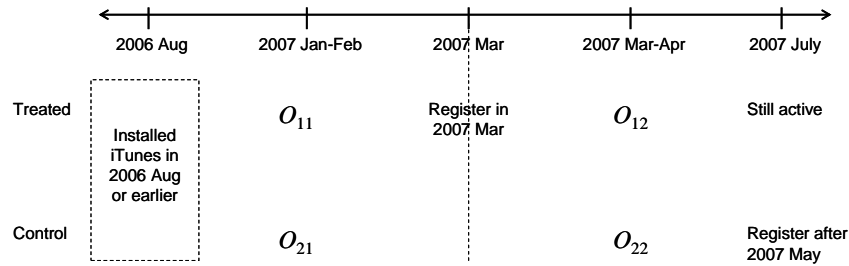
The original data comprise users who registered for Service between January – July 2007. We define the treated group as those users who registered sometime during March 2007. March is chosen because it is roughly in the middle and provides us with sufficient pre/post data. The time periods for the before-after comparison are the two-month windows January-February and March-April.<sup>5</sup> The control group is defined as users who registered for Service sometime from May on. We observe this group's Service-unaffected behavior over January-April because upon their eventual registration, sometime from May on, we extract their iTunes history files and look backward at the January-April period.

A criterion for inclusion in the study is that each user began using iTunes in August 2006 or earlier. Upon installing iTunes or buying an iPod, users often load their CD collections onto their computers. We do not want to treat loading of old music as new purchases. Thus the criterion of installing iTunes in August 2006 or earlier creates a buffer of at least four months (September-December 2006) between installing iTunes and our analysis. This is conservative because the loading of old CDs typically

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<sup>5</sup> That some users registered in late March could dampen the results' magnitude because it allows some Service-unaffected data to enter the post-recommendations period. One cannot circumvent this by centering each user's before-after data exactly on his registration date: since each user differs in this date, there would be no well-defined period for constructing the control. We are conservative and accept this tradeoff of a possible dampening of results in order to have a well defined control group.

occurs within the first month of iTunes/iPod use.<sup>6</sup>



**Figure 3.** Data Setup and Analysis Design

The second criterion for inclusion is active user status. Some users uninstall the plugin before the study’s end. So that our panel includes the same users before and after, which is required for our user-to-user before-versus-after comparisons, we adopt the criterion that users have the plugin installed for the study’s entire duration.<sup>7</sup> The implications of these data-inclusion criteria are discussed below.

The data collection has two limitations. First, assignment to the treated versus control group is not randomized. Since registration is the user’s choice, there can be a selection bias. For example, it is possible that registration is a response to changes in demand for music rather than a cause of it. A section later on sensitivity analysis shows this is unlikely. We defer a detailed discussion of this to Section 9. However, we address the selection issue by conducting our main analysis for a matched sample of treated and control users using Propensity Score Matching (PSM). PSM is a statistical matching technique for causal inference with observational data (Rosenbaum and Rubin 1983). PSM usually involves running a logistic regression for group membership to compute the probability of users belonging to the treated versus control group based on a set of observed predictors. Users in the treated group are then matched with users in the control group who have the same probability of treatment in order to control for confounding factors.<sup>8</sup> A weakness of PSM is that hidden biases may remain because matching cannot control for unobserved variables (Shadish et al. 2002; Pearl 2009). However, PSM works well with large samples and when a large number of pre-treatment covariates that are likely to influence group selection are available (Heckman et al. 1998; Shadish et al. 2002; Pearl 2009). As a result, PSM has been used as the primary technique for identification in several studies based on observational data (e.g., see Angrist 1998; Aral et al. 2009). We use one-to-one caliper matching algorithm with rank-based Mahalanobis

<sup>6</sup> The iPod/iTunes installation date is not recorded. We proxy it using the day the first song is added to each user’s library.

<sup>7</sup> Un-installation is not observed, so we proxy this by including those users’ whose plugin communicates with the Service at least once after the post-recommendations period.

<sup>8</sup> We match each control user to one treated user without replacement. Matching without replacement is essential to ensure that no user appears twice in the resulting dataset. This is important because our test statistic measures overlap in consumption across users. However, the order in which samples are matched introduces a source of variation. To ensure that our results are robust to the order in which matches are made, we also tried many different runs in which we shuffle the order of treated and control users in our sample. We also tried runs with different calipers. The results were qualitatively similar to the ones reported here. For more on matching without replacement, we refer interested readers to *Observational Studies* by Rosenbaum (2002).

distance and utilize the MatchIt package by Ho et al. (2007).

In order to create a matched sample, we first run a logistic regression for group membership against a number of pre-experiment behavioral covariates that specify users' music download behavior. These variables include *iTunes Installation Date*, *Library Size* (size of music library before start of experiment), *Downloads within 30 days of iTunes Installation*, *Average Monthly Downloads* and *Avg Change in Monthly Downloads*. We find imbalances in many covariates before matching and particularly for covariates *Library Size* and *Downloads within 30 Days of Installation*. After matching, users from the two groups are no longer significantly different on any of these dimensions. Table 1 shows the standardized difference of covariates before and after the matching in which all unacceptable imbalance (absolute value greater than 0.1) have all been rendered acceptable (below 0.1 according to Austin et al. 2007). For the sake of completeness, we also report *p*-values for difference in means *t*-test. All subsequent analysis is for the matched sample of users. We have also confirmed that our results are qualitatively similar for the original unmatched sample. Those results are available upon request.

**Table 1.** Standardized difference of the covariates

Covariates	Standardized Difference		t-test p-value	
	Before	After	Before	After
iTunes Installation Date	-0.14	0.01	<0.01	0.86
Library Size	0.23	0.07	<0.01	0.18
Downloads within 30 Days	0.23	0.09	<0.01	0.07
Avg Downloads per month	0.10	0.09	0.01	0.05
Avg Monthly change in Downloads	-0.09	-0.04	0.03	0.39
(Library Size)^2	0.07	0.04	<0.01	0.25
(Downloads within 30 Days)^2	0.15	0.07	<0.01	0.17

A  
limitation

users who

second

involves

uninstall the

plugin. Several users in the treated group uninstall the plugin before the data collection ends (before the end of period  $t = 2$ ). If the uninstallation decision is independent of music preference – for example, uninstalling the plugin to free up disk space or not liking the extra screen space occupied by the plugin – then the conclusions are unaffected because the selection is equivalent to our taking a random sample. If they are not independent, then the analysis of the non-attributing population may overstate the magnitude of the results but it will not change the direction of the results. This idea is discussed and bounded in the section on sensitivity analysis.

The resulting dataset after matching and applying the above inclusion criteria consists of 858 users each in the treated and control groups. Treated users purchased a total of 97,226 songs from 31,395 artists in the before period whereas control users purchased 106,431 songs from 32,163 artists in the before period.

## 6. RESULTS ON THE OBSERVED DATA

This section shows how the consumer network changes when recommendations are introduced. Overall, we find consumers become more similar to one another: in median degree and average distance.

### 6.1. Aggregate Analysis on the Observed Data

Using the two group design, we construct the four networks – before and after recommendations for the treated and control – and calculate the summary measures  $T(\cdot)$  on each. Then, for each summary statistic  $T$ , we calculate the changes over time  $D_1 = T_{12} - T_{11}$ ,  $D_2 = T_{22} - T_{21}$ , and the difference-in-differences estimator  $D = D_2 - D_1$ .

Table 2 shows the results. Across the columns are the two statistics: median degree, and average distance. The rows present the statistics for the treated group (row “T”) and control group (row “C”). The table’s elements show the values of  $T$  before and after recommendations. The column  $D_i$  lists the difference for each group. The column  $D$  lists the difference-in-differences estimate  $D$  and the last column  $p$  lists the  $p$ -value from a test that  $D = 0$ . To test the hypothesis that  $D = 0$ , we use the non-parametric method of permutation tests with 1000 iterations to draw the two groups. We describe the test in detail in the appendix.

**Table 2.** Summary Measures for the *Observed Data*

	Median Degree					Average Distance				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
T	850.29	845.24	-5.05	-4.47	<0.0001	0.9903	0.9842	-0.0061	-0.0052	<0.0001
C	851.49	850.91	-0.58			0.9919	0.9910	-0.0009		

The results show that under both measures, users’ purchases are more similar to one another after recommendations. First, median degree in the treated group decreases by -5.05 (i.e., the median treated user is closer to other users in the network) while the control only decreases by -0.58. The difference-in-differences  $D$  is -4.47 and it is significant ( $p < 0.001$ ). The effect is sizeable. The standard deviation of the median degree for both the treated and control groups is around 0.28 and thus the  $D$  represents a change that is almost sixteen times the standard deviation. Similarly, the average pairwise distance in treated decreases more than the average distance in the control, giving  $D = -0.0052$ . The result is statistically significant ( $p < 0.001$ ). It again represents a sizeable effect: the change is over fourteen times the standard deviation of the average distance.<sup>9</sup> Users who see recommendations become closer, whereas control

<sup>9</sup> The standard deviations are low because the cosine measure, which is based on angles between the purchase vectors, is not a sensitive measure. Given that there are thousands of artists in our dataset, most vectors are nearly orthogonal initially and most distances are above 0.95. Further, an increase in commonality by (say) 10-15 artists does not change the angle between two vectors by much. Nonetheless, the observed differences are both economically and statistically significant as highlighted above.

users do not.

## 6.2. Disaggregate Cluster Analysis on the Observed Data

The above analysis showed that users become closer after recommendations. This section asks if there are differential effects: could close users become closer but far ones farther in such a way that the aggregate result masks this? For example, users consuming Jazz music may get more Jazz recommendations and thus closer to each other but fragment away from users who consume more Classical music. If true, even though the network is more similar in aggregate, far users becoming farther would be another type of fragmentation.

This question is similar to asking whether users form tighter clusters after recommendations. If so, the aggregate effect could mask a world in which within-cluster similarity increases but between-cluster similarity decreases. We turn to cluster analysis to assess this.

We cluster users based on their artist-purchase vectors so that users with similar consumption are grouped together. We use  $k$ -means clustering method with Hartigan’s criterion (Hartigan 1975) and average silhouette width to determine the optimal number of clusters. In general, there is no “true” solution for clustering, and these are common approaches to determine number of clusters. Both Hartigan’s criterion and average silhouette width indicated that the optimal number of clusters is 5.

By comparing the within-cluster and between-cluster user distance before and after, we assess whether it is close users (those in the same cluster), far users (those in different clusters), or both close and far users that become closer after recommendations. For users in each treated before ( $O_{11}$ ) and control before ( $O_{21}$ ) group, we run  $k$ -means clustering to obtain the cluster memberships. Then we ask what happens to the distance between users who were originally in the same cluster and between users who were originally in different clusters. We report the average distance between users, which was also used in the aggregate analysis except now the measure has been separated into within-cluster and between-cluster. Median degree depends on the number of users in the network which were comparable for the treated and control groups in the aggregate analysis because both groups had 858 users. After clustering, the two groups have different number of users within any cluster and thus median degrees are no longer comparable. An alternative is to use median distance, which provides qualitatively similar results.

Table 3 presents these results. Both the within and between difference-in-difference estimate  $D$  are negative and statistically significant. Relative to the control, treated users within and across clusters are becoming similar in the music they consume. There is no evidence to suggest that within-cluster users are becoming similar while the clusters themselves are separating.

---

The unweighted network presented in Section 9.1 offers more sensitive measures and we observe that the standard deviations and the magnitude of changes are higher for that network.

**Table 3.** Summary measures for the Cluster Analysis on *Observed Data*

	Within Average Distance					Between Average Distance				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
T	0.9785	0.9809	0.0024	-0.0056	0.002	0.9937	0.9851	-0.0086	-0.0053	<0.001
C	0.9793	0.9875	0.0081			0.9951	0.9918	-0.0033		

Note that the within average distance for the treated group increases from before to after. But this is not a differential effect. We expect some chance fluctuation: users who were by chance closer revert to being farther, and users who were by chance farther revert to being closer. This is seen in the control group, where within average distance increases considerably. Many users who were by chance close to each other regressed to being farther. In summary, cluster analysis supports the result from aggregate analysis and further shows that both close and far users come closer due to the recommendation system indicating the absence of differential effects.

## 7. VOLUME EQUALIZATION

The results so far show that similarity increases under recommendations. We next explore the mechanism by decomposing the result into volume and product-mix effects.

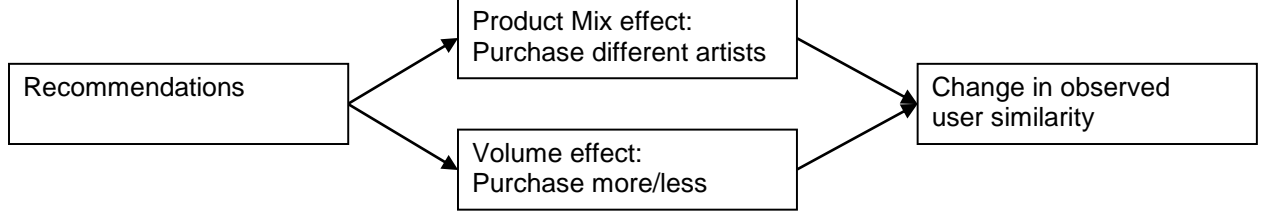
Table 4 suggests why this decomposition is needed. As the table shows, the recommender system appears to be working: users consume more after recommendations whereas control users do not. ((In fact the number of songs purchased in the control group decreases.) . The number of artists with at least one song purchased increases considerably for the treated group, indicating that users explore a wider range of artists under recommendations. Again, no such increase is seen in the control.

**Table 4.** Summary statistics for the two-groups

	Treated		Control	
	Before	After	Before	After
Users	858	858	858	858
Songs purchased	97,226	173,088	106,431	97,553
Artists with at least one purchase	31,395	56,211	32,163	28,812

These facts raise the question of whether the volume alone is responsible for creating more similarity. After all, the more consumers purchase, the more likely they are to share *some* artist in common. We thus decompose the recommender’s effects into *volume* and *product mix* components. The volume component is the portion of  $D$  due to a change in purchase volume. The product mix component is the portion of  $D$  due to changes in the assortment of artists users buy, with volume held equal. Figure 4 illustrates this, showing that recommenders can change user similarity in two ways. Both are valid ways

for recommenders to affect similarity, but we wish to distinguish them to understand the mechanism behind the overall result.



**Figure 4.** Changes in observed user similarity may have product mix and/or volume components

Until now,  $D$  was calculated on the observed data, for which volume increased after recommendations. This represented the combined product mix and volume effects. Now, we equalize purchase volume before versus after but in a way that maintains the differences in the types of music users buy before versus after. Recalculating  $D$  on the volume-equalized data then identifies the standalone product mix effect, if it is present. To equalize the volume before versus after, we use the bootstrap (Efron & Tibshirani 1986). Instead of comparing the original purchases  $O_{11}$  and  $O_{12}$ , we compare  $O_{11}$  and  $O_{12}^*$ , where  $O_{12}^*$  is sampled randomly from  $O_{12}$  and has sample size  $|O_{11}|$ . In other words, we are sampling for the empirical distribution of  $O_{12}$  and limiting the sample size to be the same as the before period. This procedure assumes the observations are i.i.d. over time, which is a common assumption in many statistical models of purchase data (e.g., latent-class multinomial models).

$$\begin{array}{ccc}
 \left( \begin{array}{cccc} 02 & 02 & 06 & 07 \\ 00 & 04 & 01 & 04 \\ 00 & 01 & 01 & 02 \end{array} \right) & \left( \begin{array}{cccc} 05 & 05 & 11 & 14 \\ 01 & 05 & 02 & 07 \\ 00 & 02 & 01 & 02 \end{array} \right) & \left( \begin{array}{cccc} 02 & 04 & 09 & 06 \\ 00 & 02 & 00 & 03 \\ 00 & 01 & 01 & 02 \end{array} \right) \\
 O_{11} & O_{12} & O_{12}^*
 \end{array}$$

**Figure 5.** Illustrating bootstrap with 3 users and 4 artists

Figure 5 illustrates the bootstrap procedure for a case with 3 users and 4 artists. Columns represent artists and rows represent purchases by users. For example, in  $O_{11}$ , user 1 purchases two songs by artists 1 and 2, six songs by artist 3 and seven songs by artist 4.  $|O_{11}|=30$  and  $|O_{12}|=55$ . 30 purchases are randomly sampled from  $O_{12}$  to generate  $O_{12}^*$ , which has  $|O_{12}^*|=30$ . By doing so, the total purchase volume is held constant across the before and after periods. For consistency, we also equalize the volume in the control group before versus after. (This is for consistency but likely unnecessary because in the control  $|O_{21}| \approx |O_{22}|$  anyway.) Last, for consistency, we equalize the volume across  $O_{11}$  and  $O_{21}$ , reducing  $|O_{21}|$  to  $|O_{11}|$  in the same manner. Thus in the volume-equalized case, we have four data sets  $O_{11}$ ,  $O_{12}^*$ ,  $O_{21}^*$ , and  $O_{22}^*$ , all with the number of purchases equal to  $|O_{11}|$ . This sampling introduces a source of variation in the results, and thus all results are averaged over repeated trials (1000 simulations).

### 7.1. Aggregate Analysis on the Volume-Equalized Data

The aggregate analysis is repeated on the volume-equalized data, and Table 5 shows the results. The same conclusion of greater similarity after recommendations emerges. However, the magnitudes are smaller, as expected because of volume equalization. For example, the difference-in-difference of the median degree is -2.48. This magnitude is smaller than on the observed (unequalized) data, which is -4.47. In short, product-mix effect accounts for roughly 55% of the difference-in-difference in median degree. Though the magnitude of the treatment effect is smaller under volume equalization, it is still significant ( $p < 0.001$ ). The average distance shows the same conclusion. The difference-in-difference measure  $D$  is negative at -0.0023, showing treated users are closer to one another than the control users and the network is “smaller.” Here again, we reject  $D = 0$  ( $p < .01$ ), providing evidence of a standalone product mix effect. Comparing tables 2 and 5, product-mix effect accounts for a little over 44% of the difference-in-difference in average distance. To summarize, when volume is held equal, purchase similarity increases after recommendations, revealing evidence of a standalone product mix effect.

**Table 5.** Summary measures for the *Volume-equalized data*

	Median Degree					Average Distance				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
T	850.29	847.09	-3.20	-2.48	<0.0001	0.9902	0.9867	-0.0034	-0.0023	0.002
C	851.63	850.91	-0.72			0.9921	0.9909	-0.0011		

## 7.2. Disaggregate Cluster Analysis on the Volume-Equalized Data

In the cluster analysis under volume-equalization, there is again no evidence of fragmentation. Table 6 shows these results. First, within-cluster average distance reduces as before ( $D = -0.0002$ ) but is not significant suggesting that after volume has been equalized, the recommender system does not have a significant effect on close users. However, far users (between average distance) become closer with  $D = -0.0029$  at  $p$ -value <0.0001. Combined with the previous results, we conclude that the recommender clearly has an effect of bringing all users together. Closer users are brought together by volume effect of the recommender while far users are brought together by both volume effect and product mix effect. Contrary to theories that recommenders keep far users far, these are the people whose similarity changes the most.

**Table 6.** Summary measures for the Cluster Analysis on *Volume-Equalized Data*

	Within Average Distance					Between Average Distance				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
T	0.9760	0.9832	0.0072	-0.0002	0.898	0.9939	0.9877	-0.0061	-0.0029	<0.0001
C	0.9792	0.9866	0.0074			0.9952	0.9920	-0.0032		

## 8. PRODUCT DISCOVERY AND EXPLORATION

What accounts for this increase in commonality? Is the recommender simply suggesting the same item to

everyone? Or is it exposing users to new types of content. The former, while demonstrating greater overall commonality, would nonetheless signal narrow media consumption and would be consistent with criticisms that these systems confine us to our information neighborhoods. In contrast, the latter would suggest that recommenders expand our horizons, help us explore and discover new types of content and to connect with other people. This, in our opinion, is the true promise of recommenders.

We examine this issue in two ways. First, we compare the changes in Gini coefficient for the two groups. The Gini is a common measure of purchase diversity, with a Gini of 0 indicating that all products have equal sales and a Gini of 1 indicating that one item generates all purchases. Higher values of Gini indicate limited diversity in the products consumed by users (see Fleder and Hosanagar (2009) for details on how to compute the coefficient). Table 7 lists the changes in the Gini coefficient for the two groups. The results indicate that there is a significant increase in purchase diversity after recommender use by treated users. One drawback of the Gini coefficient is that it is an aggregate measure of diversity and does not shed enough light on diversity at the individual level. Hence, we also compare the change in the number of unique artists consumed per user in the two groups. A *t*-test for paired observations reveals that treated users experience a significantly greater increase in artists consumed relative to the control group.<sup>10</sup> We also conduct a difference-in-difference test based on the average number of unique artists consumed by treated users versus control users (as before, the null distribution for the test is computed using a permutation test). The results are in Table 7. There is a significant increase in the number of artists consumed by treated users relative to control users, which suggests that the recommender helps them explore more artists. This is true even under volume equalization. In short, the recommender in our setting is increasing commonality in consumption, not by recommending the same items to many users but by increasing the diversity of items consumed by the users.

**Table 7. Summary measures for product exploration and discovery**

	Gini Coefficient					Number of Unique Artists				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
T	.906	.877	-0.03	-0.03	<0.0001	9709	13741	4032	4679	<0.0001
C	.905	.908	0.00			9928	9281	-647		
Volume-Equalized Data:										
T	.897	.879	-0.02	-0.02	<0.0001	9458	10488	1030	1065	0.02
C	.897	.898	0.00			9316	9281	-35		

## 9. ROBUSTNESS CHECKS

Below, we discuss the robustness of our results to other network definitions. We also examine the impact of two data limitations: non-randomized group assignment and uninstallation of the plugin.

### 9.1. Results for Other Network Definition

<sup>10</sup>  $p\text{-value} = 8.05 \times 10^{-6}$

The analysis thus far used a weighted network of users based on the cosine distance. In this section, we test sensitivity to other network definitions. This section presents one example, an unweighted network. In the online appendix, we present results for other network definitions, weighted and unweighted. All of the networks tested yield nearly the same conclusions. We present the aggregate analysis on both observed and volume-equalized data using the unweighted network.

Our unweighted network has a simple construction. Within a given group and time period, users  $x$  and  $y$  have an edge between them if they purchase at least one artist in common.

#### Unweighted

$$(G_{it})_{xy} \equiv \begin{cases} 1, & \text{if users } x \text{ and } y \text{ have } \geq 1 \text{ artist in common } ((O_{it})_x \bullet (O_{it})_y \geq 1) \\ \text{Unconnected,} & \text{otherwise} \end{cases}$$

This is an unweighted network in which any edge, if it exists, has weight 1. The  $\bullet$  symbol indicates the vector dot product, showing how this definition might be generalized to other similarity functions.

In an unweighted network, the summary statistics are slightly different. With the network  $G(O_{it})$  defined, we next define summary statistics of the network's properties,  $T(G(O_{it}))$ .

Measure	$T(G(O_{it})) =$
Density	$\frac{1}{n C_2} \sum_{x=1}^n \sum_{y < x} I(G_{xy} = 1)$
Median Degree	$\text{Median}\{d_x\}_{x=1}^n$
Path Length	$\frac{1}{n C_2} \sum_{x=1}^n \sum_{y < x} \text{Shortest Distance}(x, y)$

*Density.* The density is the fraction of edges that exist out of the total number of edges possible. Higher density means users have more connections among them.

*Median Degree.* The median degree is the number of connections to other users that the typical (median) user has. Unlike the definition from the weighted network, increase in median degree means higher similarity among users.

*Path Length.* The path length is the shortest distance between any two users, averaged over all users in the network. If users  $x$  and  $y$  are connected, the shortest distance is 1, the edge between them. Otherwise, the path is through other users. The shorter this distance, the “smaller” the network is said to be, using the terminology of Watts & Strogatz (1998), who popularized the study of “small world” networks. Mathematically, the shortest distance between users does not have a closed form expression, but it can be computed using Dijkstra’s algorithm or the Floyd-Warshall algorithm (Papadimitriou & Steiglitz 1998).

We present the aggregate analysis equivalent to that in Section 6.1 for observed data and in 7.1 for volume-equalized data. Table 8 shows the results for observed data using unweighted network. The results show that on all three measures, users' purchases are more similar to one another after recommendations. First, the treated network becomes denser, showing that users have more connections among themselves. Before recommendations, 24% of the edges are filled in, and after 46% are present, yielding  $D_1 = 22\%$ . This is a large increase in density. Over the same period, the control has no noticeable change and  $D_2 \approx -1\%$ . The difference-in-differences estimate is  $D = 23\% > 0$ , indicating that the treated network does become more dense relative to the control. This difference is significant, as the hypothesis  $D = 0$  is rejected ( $p < 0.01$ ). The median degree increases,  $D > 0$ , indicating that the typical user has more connections to others. Similarly, the path length decreases,  $D < 0$ , indicating that on average users are fewer hops away from one another. All of the results are significant ( $p < 0.01$ ).<sup>11</sup>

**Table 8.** Summary Measures for the Unweighted Network – *Observed Data*

	Density					Median Degree					Path Length				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
T	24%	46%	22%	23%	<0.01	157	361	204	207	<0.01	1.79	1.53	-0.26	-0.25	<0.01
C	20%	20%	-1%			126	123	-3			1.86	1.86	0.00		

Table 9 shows the result for the volume-equalized data. The same conclusion of greater similarity after recommendations emerges. Again, the magnitudes are smaller, as expected because of volume equalization. For example, the difference-in-difference,  $D$ , in network density decreases from 23% to 10%, which implies that product-mix effect accounts for nearly 44% of the original  $D$  and the volume effect accounts for the remaining 56%. The other measures show the same conclusions: the median degree increases, showing users have more connections to one another, and the average path length decreases, showing that users are closer to one another and the network is “smaller.” In every case we reject  $D = 0$  ( $p < .01$ ), providing evidence of a standalone product mix effect. Product-mix effect explains roughly 47% of the difference-in-difference in median degree and 48% of that in path length.

**Table 9.** Summary measures for the unweighted network – *Volume-equalized data*

	Density					Median Degree					Path Length				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$

<sup>11</sup> All the networks have one large, connected component containing nearly all users with few unconnected users outside it. Thus the density, degree, and path lengths are not biased due to changes in the size of the main component.

T	24%	35%	0.11	10%	<0.01	150	256	106	95	<0.01	1.80	1.65	-0.15	-0.12	<0.01
C	19%	20%	0.01			112	123	11			1.88	1.86	-0.03		

Clustering analysis on the unweighted network yields qualitatively the same results as that for the weighted network. We do not replicate the results below, and they are available in the online appendix. Instead, we share results from complementary analysis that help illustrate the impact of the treatment on near versus far users for both the groups. Table 10 presents the path length between all  $nC_2$  user pairs in the unweighted network. The horizontal axis is the number of hops before recommendations, and the vertical axis is the number of hops after recommendations. The values in the table are the percent of user-pairs falling in each cell. For example, 8.01% of the treated users were one hop away before recommendations and 2 hops away after recommendations. User-pairs becoming farther lie above the diagonal, while user-pairs becoming closer lie below it. A distance of infinity ( $\infty$ ) means there is no path between the given two users.<sup>12</sup>

**Table 10.** Path Lengths between all user pairs – *Observed data*.  
Entries represent the percentage of all  $nC_2$  user-pairs.

Treated					Control				
# Hops After	$\infty$	3	2	1	# Hops After	$\infty$	3	2	1
	0.12	0.02	8.01	16.07		0.36	0.26	11.8	7.72
	1.03	0.14	37.94	28.41		4.08	3.76	47.62	11.1
	0.20	0.10	2.71	0.83		0.84	1.12	4.02	0.44
	0.31	0.09	3.03	1.00		2.09	1.04	3.20	0.37
	1	2	3	$\infty$		1	2	3	$\infty$
	# Hops Before					# Hops Before			
Becoming closer (below diagonal)	36.0				Becoming closer (below diagonal)	20.2			
Becoming farther (above diagonal)	9.5				Becoming farther (above diagonal)	21.1			
No change (on diagonal)	54.4				No change (on diagonal)	58.5			

The control group appears stable (right side), as it has roughly equal weight above and below the diagonal. In contrast, the treated group (left side) has more user-pairs becoming closer (36.07% weight below the diagonal) than there are becoming farther (9.51% weight above the diagonal). This is consistent with the findings above. Second, the increase in similarity appears uniform: all types of users become closer to one another. Users who were close became closer, *and* users who were initially far became

<sup>12</sup> In Table 10, a very small number of pairs are four or five hops away. This number is so small ( $< 0.04\%$ ) that for clarity we omit them from the presentation (but not the analysis) to avoid rows and columns of nearly all zeros.

closer too. There does not appear to be evidence of a differential effect.

As before, some treated users do grow farther, but this is not a differential effect. In the control group, 11.8% went from 1 to 2 hops while 11.1% went from 2 to 1 hops. This level of mixing is roughly equal. In the treated group, some pairs do become farther – 8.01% go from 1 to 2 hops – but many more become closer – as 28.41%, went from 2 to 1 hops. To summarize, the trend toward greater similarity exists at all initial path lengths, independent of whether users were initially close or far away.

Under volume-equalization, there is again no evidence of differential effects. Table 11 shows these results. First, the aggregate effect toward similarity in the treated group is evident: there are more users becoming closer than there are becoming farther (28.71% weight below the diagonal versus 14.59% weight above it). The control group is roughly balanced. This is consistent with the aggregate findings. The magnitude is again smaller, as expected, because volume equalization dampens the effect. Second, the increase in similarity appears uniform: users who were close became closer, and users who were initially far became closer too.

**Table 11.** Path lengths between all user pairs – *Volume-equalized data*.

Entries represent the percentage of all  $nC_2$  user-pairs.

		Treated						Control			
# Hops After	$\infty$	0.25	2.34	0.59	0.39	# Hops After	$\infty$	0.43	3.80	0.85	2.39
	3	0.07	0.63	0.14	0.35		3	0.4	3.81	1.15	1.28
	2	10.71	43.83	2.74	3.12		2	10.8	47.32	4.10	4.09
	1	12.33	21.04	0.62	0.84		1	7.14	11.47	0.49	0.50
		1	2	3	$\infty$			1	2	3	$\infty$
		# Hops Before						# Hops Before			
Becoming closer (below diagonal)		28.7				Becoming closer (below diagonal)		21.9			
Becoming farther (above diagonal)		14.5				Becoming farther (above diagonal)		20.9			
No change (on diagonal)		56.6				No change (on diagonal)		58.0			

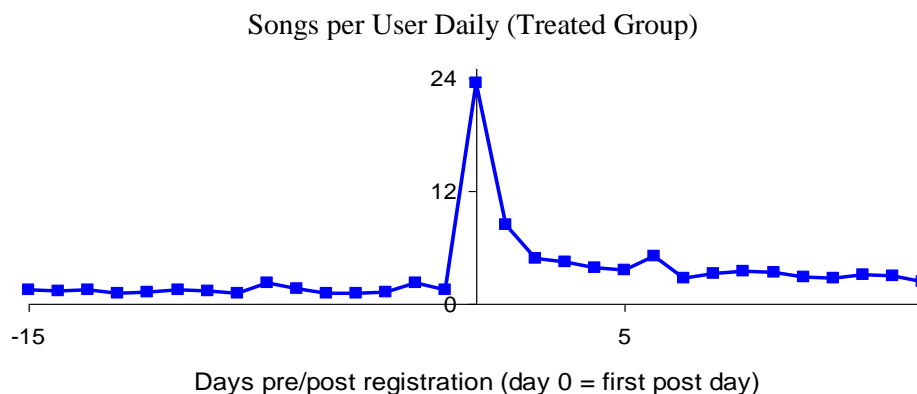
## 9.2. User Registration Decision

One limitation of the data collection is that assignment to the treated versus control group is not randomized. Registration is the user's choice, so the analysis cannot account for selection on unobservables. For example, both the registration decision as well as our observed changes in purchases similarity may be driven by changes in users preferences around the time of registration. We addressed this issue by conducting all analysis on a matched sample of users, as is commonly done in the literature.

In this section, we exploit some unique features of our dataset to conduct additional tests that show that it is unlikely that our results are due to a selection bias. We begin by presenting several arguments for why we believe this is unlikely. Next, we share results from a more formal investigation of selection bias along the following lines: (a) Ruling out a time trend among treated group, (b) Verifying impact of treatment on the control group.

We first note that both the treated and control users in this study are both eventual users of the recommender system. Thus, the selection issue is not as acute as is typical in many observational studies in which control users do not select the treatment. In our setting, the control users also select the treatment and do so only a few weeks later. This, by itself, ensures a high level of similarity between the two groups. Further, we attribute the small differences in adoption timing between the two groups primarily to diffusion of product awareness as opposed to fundamentally different demand preferences. This is because Service was a new technology at the time of data collection and the very first iTunes plugin of its kind. Registration may thus be reasonably seen as a response to a change in supply rather than a change in consumers' own demand. And differences in registration timing among early users may similarly be viewed as arising primarily due to spread of product awareness. This line of reasoning is the same as Waldfogel and Chen's study (2006) of how sales at unbranded retailers are affected by the introduction of comparison shopping engines on the web, which were at the time a new technology.

To test this idea, Figure 6 shows the median number of songs treated users add to their libraries in the days before and after registration. The data are centered around each user's registration date. The figure shows that the change in behavior is sharp near registration and not part of a growing trend starting weeks before. We now test the robustness of our results more formally.



**Figure 6.** Daily songs added per user (median) centered on each user's registration date. Day 0 represents the time immediately after registration.

Ruling out a time trend among treated users: One possibility is that the treated group had been experiencing changes in preferences in the days preceding registration and our results merely reflect these

time trends rather than the impact of the recommender. Figure 6 suggests this is unlikely. We test this more formally by conducting a Difference-in-Difference (DiD) test of purchase similarity in multiple pre-treatment periods (Meyer 1995). The “before” period for this test is defined as January 2007 and the “after” period is defined as February 2007. Note that both groups had not been exposed to the recommender system during this entire timeframe. However, if the treated users were experiencing change in preferences over time, then we expect these changes to show up in the DiD test. Table 12 shows that there are no significant changes in median degree and average distance for the treated users relative to the control users. Thus, we can rule out the possibility that our results reflect a time trend of increased purchased similarity among the treated users.

**Table 12. Summary measures for pre-registration time periods**

	Median Degree					Average Distance				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
T	655.70	656.00	0.3037	-0.3404	0.4020	0.9935	0.9933	-0.0002	-0.0007	0.3480
C	656.06	656.70	0.6441			0.9942	0.9947	0.0005		
Volume-Equalized Data:										
T	655.99	655.99	0	-0.4650	0.2360	0.9940	0.9933	-0.0007	-0.0009	0.2320
C	656.38	656.85	0.4650			0.9948	0.9950	0.0002		

Effect of treatment on control users: A unique aspect of our dataset is that the control users also registered for the recommender a few weeks after the treated users. If these control users do not demonstrate a similar change in purchase similarity upon registration, then it might suggest that the recommender system may not be driving the observed changes and that the treated users in our study are fundamentally different from the control users. This is similar to the analysis by Gruber (1994) in which a later federal mandate on maternity benefits (the treatment) resulted in some states that had not previously mandated such benefits (the original control states) to now be subject to the treatment. To evaluate the effect of the treatment on our original control group, we divide our control users into two groups. The first group, G1, registered for the recommender in May 2007 and the second group, G2, registered in July or August 2007. We consider March and April as the “before” period and May and June as the “after” period. Note that G1 users are exposed to recommendations in the after period whereas G2 users are unexposed throughout. Table 13 shows that we observe a significant decrease in median degree and average distance for G1 users relative to G2 users. Thus control users also experience an increase in purchase similarity when they are exposed to the recommender.

**Table 13. Summary measures for the original control users**

	Median Degree					Average Distance				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
G1	809.36	805.34	-4.0190	-3.8929	<0.0001	0.9912	0.9860	-0.0052	-0.0048	<0.0001
G2	809.80	809.68	-0.1260			0.9912	0.9909	-0.0003		

	Volume-Equalized Data:									
G1	809.37	806.63	-2.7434	-2.4779	0.002	0.9912	0.9879	-0.0033	-0.0028	0.002
G2	810.08	809.81	-0.2654			0.9917	0.9913	-0.0005		

### 9.3. Attrition

The second data limitation is attrition. About half of the users in the treated group uninstall the plugin before the data collection ends. The above analysis, as discussed, only considers those users who have Service installed for the study’s duration. Although attrition is a common issue in all observational data and is not unique to our setup, we nonetheless provide a brief discussion of its impact.

The implication of attrition is that we may overstate the magnitude of the results although not their direction. This conclusion requires the assumption that uninstallers return to pre-treatment behavior and resemble the control group. To illustrate how attrition affects the magnitude, we can “average” the treated users who complete the study with control users as proxies for the drop-outs. From the previous results, we saw that the treated group’s similarity increases and the control’s shows almost no change, so “averaging” the results dampens the magnitude but not sign. A key question, however, is whether the increase in commonality continues to be significant even after the dampening. We test this below.

To estimate the effect of attrition, suppose the treated group originally has  $n$  users and  $\lambda n$  uninstall the service ( $0 < \lambda < 1$ ). We observe the  $(1 - \lambda)n$  users who remain with Service. Under the assumption that the drop-outs resemble the control, we can approximate the original treated group using all  $(1 - \lambda)n$  treated users and  $\lambda n$  control users. We refer to this group as *Composite*. We simulate a Composite group by randomly drawing users without replacement from the control group to replace treated users who uninstall. In our data,  $\lambda \approx 0.5$ , and Table 14 presents the results for the weighted network. For the observed data, the results show that there is a significant increase in purchase similarity among Composite users even after accounting for attrition. For volume-equalized data, statistical significance is lost but the direction stays the same for the weighted network. Two comments are in order. First, in several environments, firms deploy recommenders and users do not have an option to “turn off” personalization and thus the dampening of magnitude for Composite users may not occur. Second, the weighted network measure is not very sensitive. Even if a user adds many new artists in the after period, this may not significantly change the angle between the user’s purchase vector and that of another user because of the large number of artists in the dataset (and the resulting high-dimensional space). One way to test this idea is to repeat the analysis for an unweighted network. Table 15 shows that both observed data and volume-equalized data are still significant at  $\lambda=0.5$ . Finally, we repeat the simulations for different values of  $\lambda$  and find that significance is lost for the weighted network (at  $p=0.05$ ) at  $\lambda=0.67$  (i.e., if 67% of treated users had uninstalled the service) for the observed data and  $\lambda=0.42$  for the volume-equalized data.

**Table 14.** Summary Measures for *Composite* Weighted Network ( $\lambda = 0.5$ )

Median Degree						Average Distance				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
T	850.9983	848.329	-2.6693	-2.0891	<0.0001	0.9912	0.9878	-0.0034	-0.0024	<0.0001
C	851.4957	850.9155	-0.5802			0.9919	0.991	-0.0009		
Volume-Equalized Data:										
T	851.1542	849.0389	-2.1153	-1.3949	0.06	0.9914	0.9888	-0.0026	-0.0015	0.116
C	851.6458	850.9254	-0.7204			0.9921	0.991	-0.0011		

**Table 15.** Summary Measures for *Composite* Unweighted Network ( $\lambda = 0.5$ )

Density						Median Degree					Path Length				
	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$	Before	After	$D_i$	$D$	$p$
T	22%	32%	10%	11%	<0.01	148	252	104	107	<0.01	1.84	1.70	-0.14	-0.14	<0.01
C	20%	20%	-1%			126	123	-3			1.86	1.86	0.00		
Volume Equalized Data:															
T	20%	26%	6%	4%	<0.01	143	206	63	47	<0.01	1.84	1.75	-0.09	-0.06	<0.01
C	17%	19%	2%			118	134	15			1.88	1.86	-0.02		

## 10. RELATIONSHIP TO SERVICE’S RECOMMENDER SYSTEM

This section relates these findings to the recommendation system in use at Service. We believe similarity increases post-recommendations because Service makes users’ choice sets more similar than if users were not members of the recommendation service. This appears true for both components of Service, the plugin and the website.

With the plugin, recommendations are based on the artist a user is currently listening to. When two people listen to the same artist, they receive the same list of recommendations. Because of this, users having the same artist are more likely to see the same recommendations and thus more likely to purchase another common item. In terms of the unweighted network from Section 9.1, users who are 1 hop away in the treated group should be more likely to remain 1 hop away than control users: Table 16 supports this. Treated users 1 hop away are 67% likely to remain 1 hop away afterward, whereas 1 hop away control users are only 38% likely to remain at 1 hop.<sup>13</sup> Seeing the same recommendations maintains the 1-hop position among treated users, whereas there is no such force for control users. Why do users not connected beforehand ( $k \geq 2$ ) become closer? Such users do not own a common artist from which identical recommendations can be generated. Recall that Service provides a *list* of recommended artists in its plugin. When a  $k \geq 2$  pair of users listens to related but different artists, their recommended lists can still include the same recommended artist. If both buy songs by this artist, the users now have a purchase in common. In this manner, treated  $k \geq 2$  users should be more likely to connect than control users. As

<sup>13</sup> The probabilities are approximated as the fraction of user pairs transitioning from  $k$  to 1 hops, and the data come from Table 10 and Table 11.

such, if this is the mechanism by which Service affects  $k \geq 2$  users, we would expect this effect to be greater for  $k = 2$  users than  $k = 3$  and in turn  $k = \infty$  users. To test this idea, one observes again in Table 16 that Prob(1 hop away after  $| k$  hops away before) does show a primarily decreasing trend.

**Table 16.** Probability(User pair is 1 hop away after  $| k$  hops away before).

Initial hops $k =$	Treated				Control			
	1	2	3	$\infty$	1	2	3	$\infty$
Observed data	0.67	0.41	0.23	0.24	0.38	0.16	0.06	0.05
Volume equalized	0.44	0.24	0.13	0.12	0.29	0.13	0.05	0.04

At Service’s website, a similar phenomenon creates co-purchases among users. When one examines another user’s play history, those are songs the other user already owns. Thus any purchase of those songs creates a co-purchase. In turn, more co-purchases results in an increase in similarity.

Finally, without variation in the components’ design, one might argue that Service could design a perverse recommender to achieve any end it wanted, similarity or fragmentation. We do not believe we are observing this perverse case. Service’s algorithm was designed to satisfy users and not for an explicit goal of creating or reducing fragmentation. Second, we believe Service’s design is somewhat typical for the industry: a content based algorithm where songs in the same sub-genre are recommended; a collaborative algorithm where songs co-purchased are recommended; and a website where one can browse other users’ profiles, as is common at many social networking sites. A large factorial design testing alternative designs for each component would certainly be desirable, and we hope future work will contribute to this.

## 11. CONCLUSIONS

Much of our time spent online – whether reading news, listening to music, or purchasing books – is guided invisibly by recommendation algorithms. Despite the many millions of hours guided by them and papers proposing new algorithms, we know much less about how they affect the market and society.

This paper asked whether recommender systems fragment versus homogenize users. Using data from the music industry, we found that users have more in common after recommendations, as measured by purchase similarity. The increase in commonality occurred for two reasons: the product mix effect, in which users shifted their purchases toward more similar items, and the volume effect, in which users simply bought more under recommendations, increasing the likelihood of co-purchases with others. Each effect contributed roughly equally to the overall result. Further, when users are clustered by taste, both the within-cluster and between-cluster distances shrink after recommendations. Personalization thus appears to be a tool that helps users widen their interests and create commonality.

Regarding policy implications, many have criticized recommenders fearing that recommenders

will fragment the online population. We agree that excessive fragmentation could be undesirable. However, our data show that recommenders appear to create commonality, not fragmentation. In the absence of such effects, there is no cause, based on this study, to modify the architecture of e-commerce or the web.

Regarding business, the study provides a window onto the ongoing trend of targeted marketing. Our study demonstrates that recommender systems can drive a significant increase in purchase volume and may further alter the mix of products users buy.

The results suggest several areas for future work. The first area is studying additional recommender technologies. This paper studied a major, online recommender system. We hope future work will look at other designs and gradually catalog their effects. Second, one could study domains other than music (e.g. news, books, fashion). The manner in which users respond to news or fashion recommendations may differ from the manner in which they respond to music recommendations. The policy implications for news would be especially important. Finally, a third area is relating recommender design choices to commonality. For example, it is possible that some artists are boundary-spanners (e.g. Elvis could be classified under both rock and country). Recommenders that explicitly promote boundary-spanning artists could play a key role in driving exploration and commonality. Future designs may want to consider this.

In *The Big Sort*, Bishop (2008) shows how over the last thirty years Americans have sorted themselves into like-minded neighborhoods. This paper asks a similar question about the virtual space of the web. While many predict these systems will further a trend of fragmentation, the evidence for the industry and firm studied here is to the contrary. As this is the first empirical study on the topic, we look forward to the perspective thirty more years will provide.

## APPENDIX I. SIGNIFICANCE TESTING

The hypotheses tested in the aggregate analysis have the form

$$H_0: \mu \equiv E[D] = 0$$

$$H_a: \mu \equiv E[D] \neq 0$$

where  $D \equiv (T_{12} - T_{11}) - (T_{22} - T_{21})$  and  $\mu \equiv E[D]$ . This is a statistical test of the null hypothesis that purchases are distributed the same in the treated group and in the control group. The use of such test statistics is facilitated by permutation tests which allow us to calculate a null distribution for any test statistic. Statistical theory says that under the null hypothesis of equal distributions of purchase records (and conditional on the observed purchase records), all relabelings of the records as 'Treated' and 'Control' are equally likely. We obtain a null distribution and hence a  $p$ -value for  $D$  by repeatedly relabeling the

purchase records, reconstructing the networks, recalculating  $D$ , and tallying the fraction of times these 're-labeled' values of  $D$  exceed the observed value of  $D$ . Enumerating all relabelings is not usually possible computationally, which is why one resorts to sampling a feasible number of relabelings that yields an approximate permutation  $p$ -value for  $D$ . Further details on the theory of permutation tests can be found in the appendix to Good (1994). In our study, we used 1000 iteration for permutation tests.

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